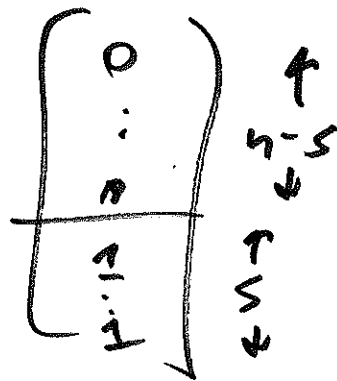
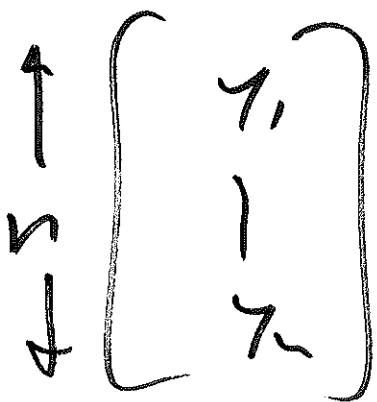


prior  
sample  
size  
 $n_0 = (\alpha + \beta)$

known info ①

mean  $\frac{\alpha}{\alpha + \beta} = E(\theta | \text{Beta}(\alpha, \beta) B)$



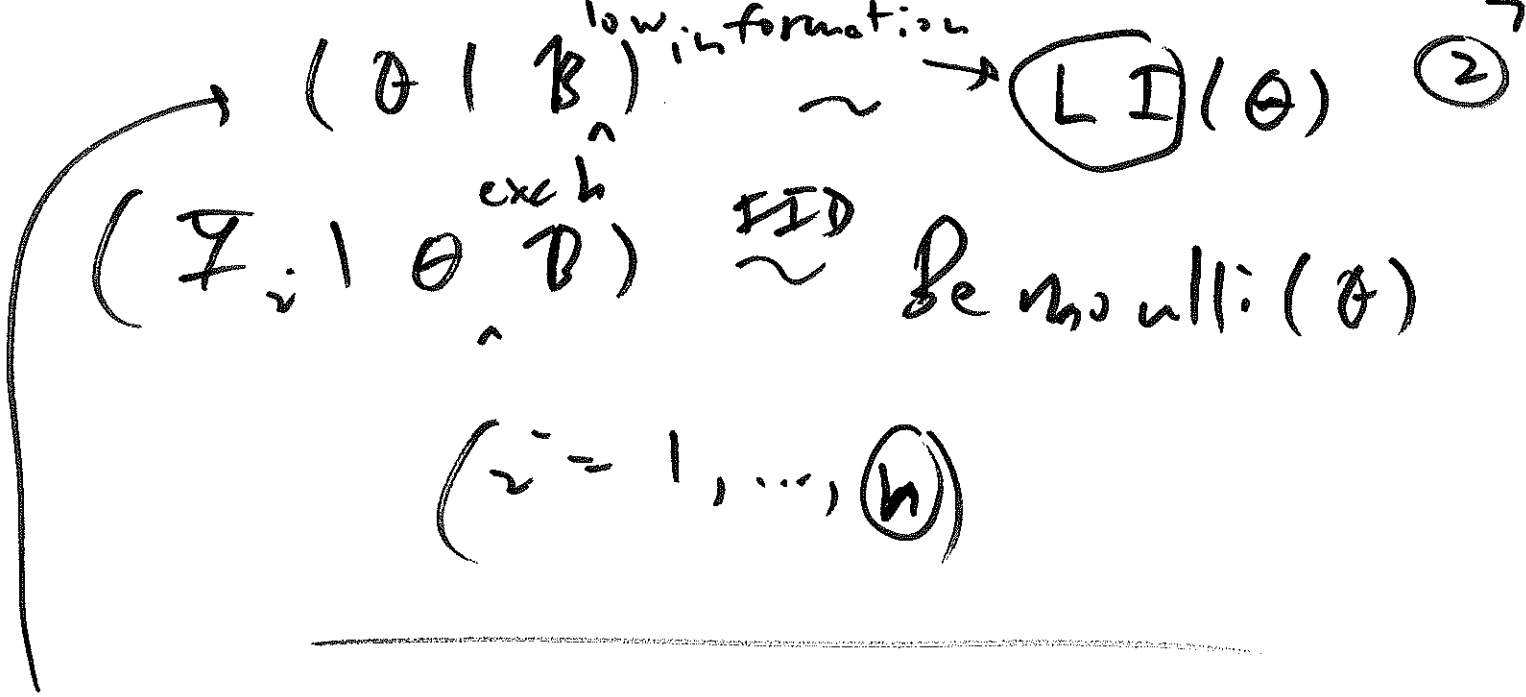
data  
info.

data  
sample  
size  
n

mean  $\bar{y} = \frac{s}{n} = \frac{s}{s + (n-s)}$

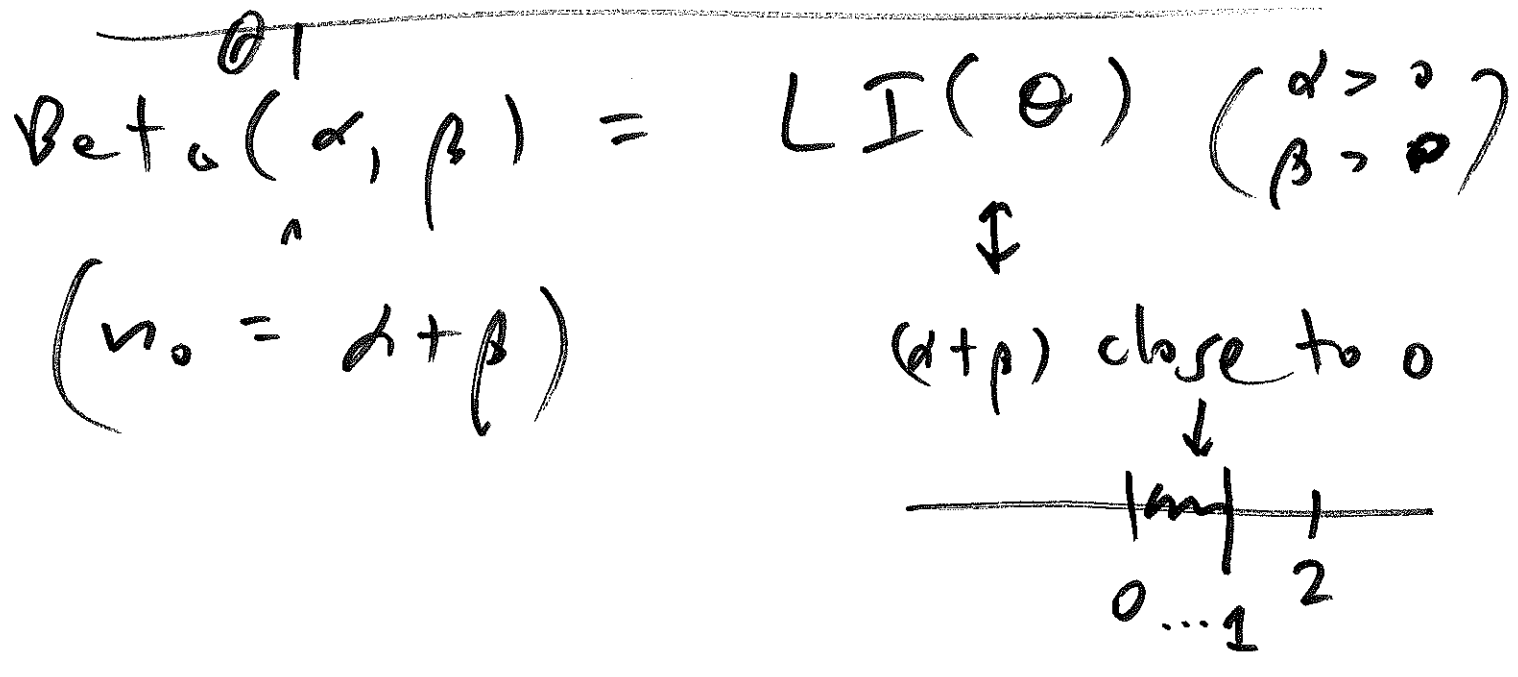
$$\frac{\alpha + s}{\alpha + \cancel{s} + (\beta + n \cancel{s})} = \frac{\binom{\alpha + \beta}{w} \binom{\alpha}{\alpha + \beta + n} + \binom{n}{1-w} \frac{s}{n}}{\binom{\alpha + \beta + n}{\alpha + \beta + n} \text{ (prior mean)}} \quad \text{data mean}$$

$\alpha + \beta$  (prior mean)       $0 \leq w \leq 1$

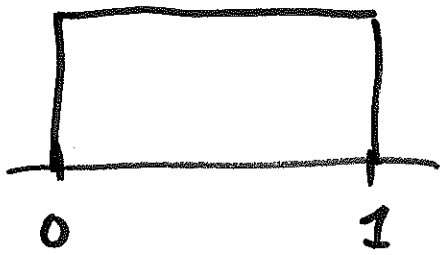


if context implies not much known about  $\theta$  external to  $\mathcal{F}$

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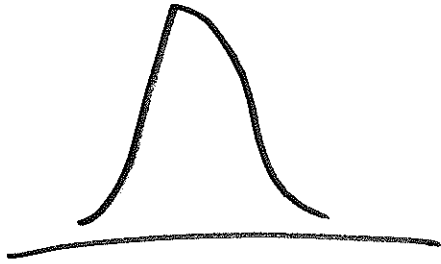
Beta(1,1) = uniform(0,1) ③



$p(\theta | \text{Beta}(\alpha, \beta) \mathcal{B})$

LI( $\theta$ ) prior

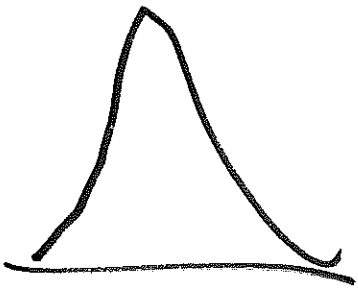
prior sample size = 2



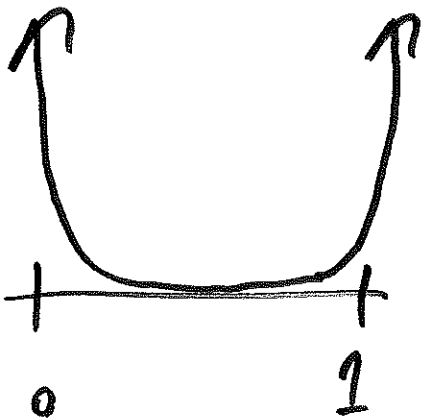
$Q(\theta | \text{exch } \mathcal{B})$

data sample size

lik.  
 $n = 403$



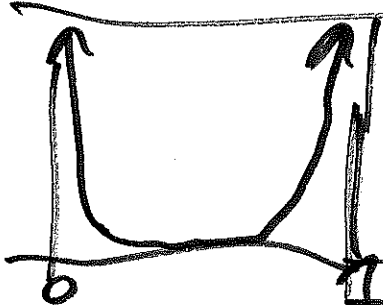
$p(\theta | \text{exch } \mathcal{B}) \sim \text{Beta}(\alpha, \beta)$  post.



Beta( $\frac{1}{2}, \frac{1}{2}$ )

prior sample size 1

(1950)  
(Jeffreys prior)



Beta( $\alpha, 0$ )

(Haldane prior) (1935)

$$\int_0^1 \text{Beta}(\theta, \theta) d\theta = +\infty$$

such a dist. is called improper

&  $\text{Beta}(\theta, \theta)$  prior with  
 IID Bernoulli lik.  $f^h$  is  
 an example of the use  
 of an improper prior

useful Bayesian  
 advice

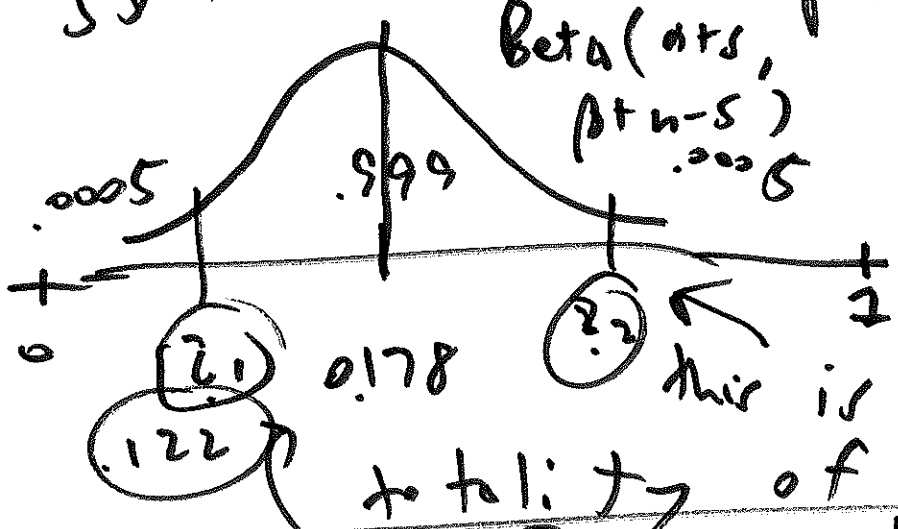
don't use  
 improper priors

V.L.W  
 = IC

p.m.s.s. =  $2\epsilon$  for  $\epsilon > 0$  & near 0  
 almost improper  $\text{Beta}(\epsilon, \epsilon)$

SD 0.0184

post. dist



$p(\theta | Z)$  (exch)  $Beta(\alpha, \beta)$   
 4.5 ↓  
 25.5 ↑

post. mean = 0.178

data mean (F)  
~~0.180~~  
 0.179

post. SD 0.0184  
 $n_{post} = \frac{433}{\alpha + \beta + 4}$

$\vec{SE}(\bar{y}) = 0.0192$   
 $n = 402$

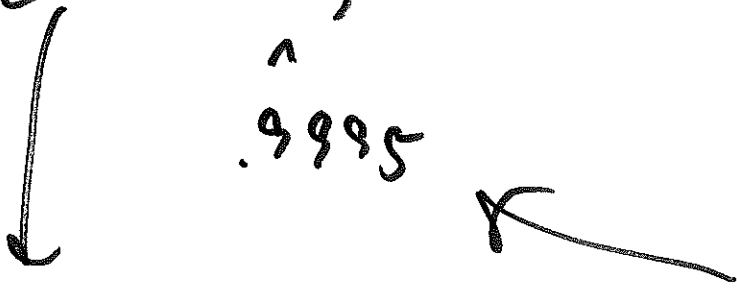
99.9% credible interval  
 (0.122, 0.242)

99.9% CI  
 (0.116, 0.241)

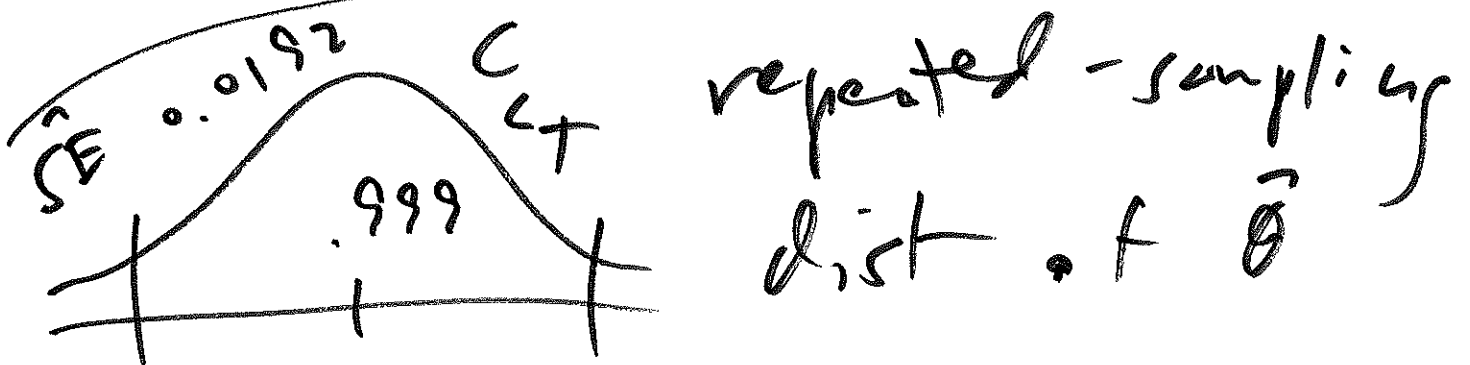
0.999 = P\_B ( ?\_1 ≤ θ ≤ ?\_2 )

beta( 4.5 + 72, 25.5 + 403 - 72 )

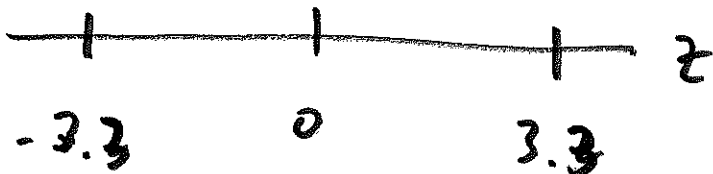
.9995



.122



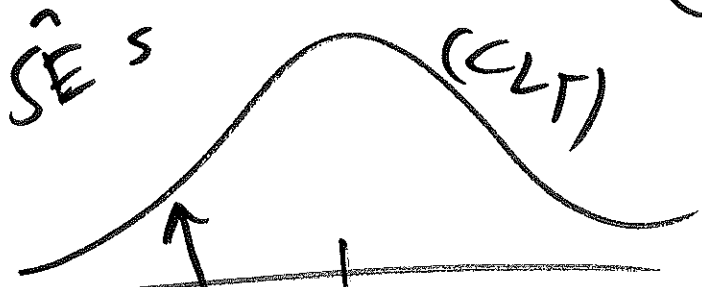
theta - 3.3 sigma\_hat\_E, theta, theta + 3.3 sigma\_hat\_E (n = 403)



gamma\_m(.9995)

99.9% CI for theta: theta-hat ± 3.3 sigma\_hat\_E(theta-hat)

(F)



$$c_1 e^{-c_2 (\hat{\theta}_{MLE} - \theta)^2} \quad \theta \text{ (fixed)}$$

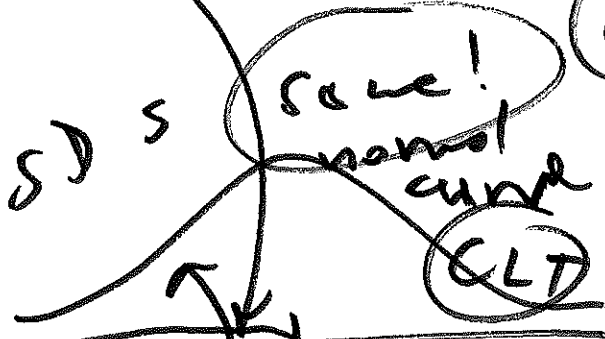
normal curve for  $\hat{\theta}_{MLE}$  in RS, centered at  $\theta$

(7)

repeated-sampling dist. of  $\hat{\theta}_{MLE}$  (R.V.)

$n$  large

(B)



same! normal curve

$$c_1 e^{-c_2 (\theta - \hat{\theta}_{MLE})^2} \quad \theta \text{ (fixed)}$$

normal curve for  $\theta$ , centered at  $\hat{\theta}_{MLE}$

post. dist. for R.V.  $\rightarrow \theta$  given  $Y$  & LI prior,  $n$  large

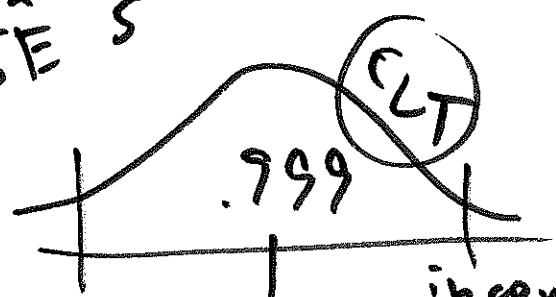
Laplace 1800

B-v-M

Bernstein-von Mises theorem: with large  $n$  & LI prior,

frequentist inf.  $\hat{=}$  Bayesian inf.

$\hat{SE}_s$



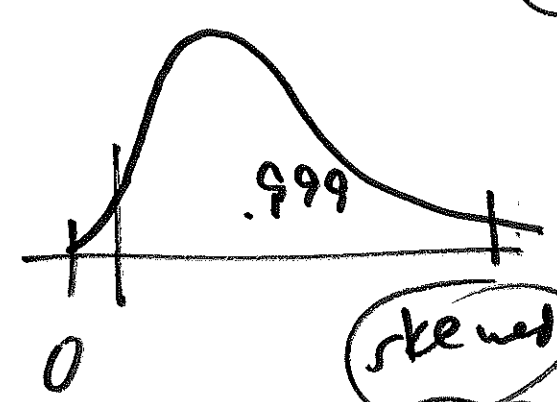
$\hat{\lambda}$  (incorrectly symmetric) may

(F) (this week's disc. sec.) (8)

RS dist of MLE,  $\hat{\lambda}_{MLE}$   $\hookrightarrow$  small

Simple versions of Neyman & Fisher pretend CLT always holds, even for small  $n$

(B)



skewed  
right

post dist for  $\lambda$  given  $\lambda$  (small)  $\uparrow$  14  
& LI prior



MLE story easier than Bayes <sup>(9)</sup>

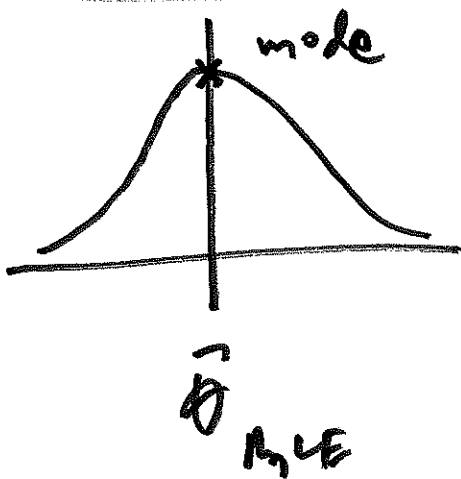
differentiation

(easier)

integration

(harder)

when B-V M holds, can  
therefore use freq. (MLE)  
as good approx to Bayes



$$l(\theta | \mathcal{Z} \text{ excl. } \mathcal{B})$$

B-V M

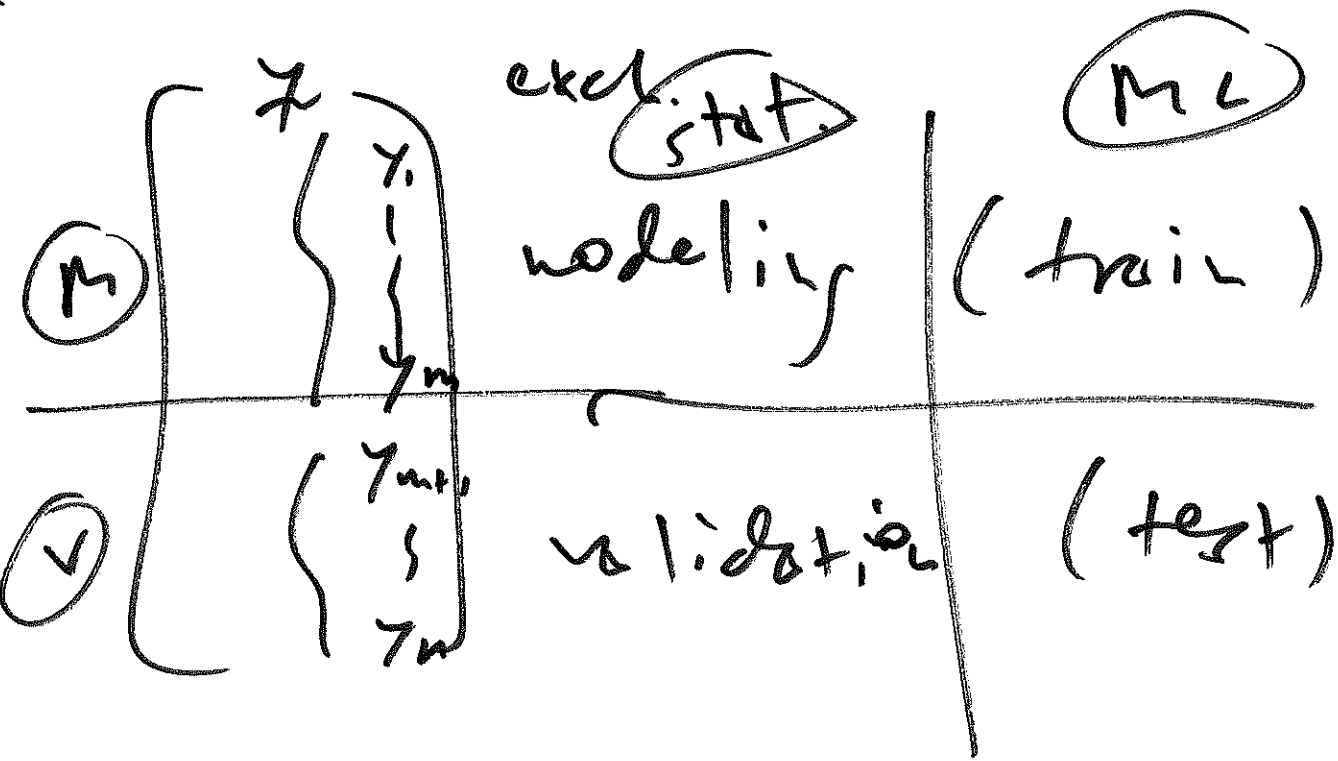
$$\approx p(\theta | \mathcal{Z} \text{ excl. } \mathcal{B})$$

$\approx$  normal

$\approx$  post. median

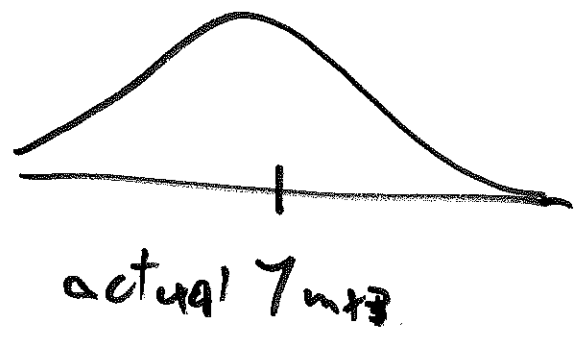
$\approx$  post. mean  $\leftarrow$  Bayesian point est.

large



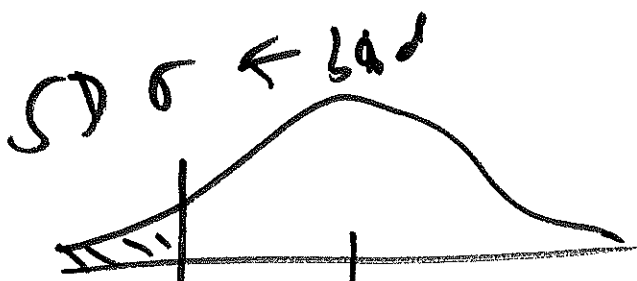
build model on  $M$  data, check its goodness by forcing it to predict  $V$  data :

2-way cross-validation



$$p(\underline{y_{m+1}} \mid y_1 - y_m \text{ model})$$

how compare a data value with a pred. list for it?



25%

$\theta$

$\uparrow$   
be  $H_0$